

A unified framework to quantify demographic buffering in natural populations

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Abstract (143/150 words)

The Demographic Buffering Hypothesis (DBH) predicts that natural selection reduces the temporal fluctuations in demographic processes (such as survival, development, and reproduction), due to their negative impacts on population dynamics. However, a comprehensive approach that allows for the examination of demographic buffering patterns across multiple species is still lacking. Here, we propose a three-step framework aimed at quantifying demographic buffering. Firstly, we categorize species along a continuum of variance based on the sums of stochastic elasticities. Secondly, we examine the linear selection gradients, followed by the examination of nonlinear selection gradients as the third step. With these three steps, our framework overcomes existing limitations of conventional approaches to quantify demographic buffering, allows for multi-species comparisons, and offers insight into the evolutionary forces that shape demographic buffering. We apply this framework to mammal species and discuss both the advantages and potential of our framework.

Environmental stochasticity plays a pivotal role in shaping organisms' life histories (Bonsall & Klug 2011). Nonetheless, how organisms will cope with the increasing variation in environmental conditions expected under climate change (Boyce *et al.* 2006; Morris *et al.* 2008) is one of the most intriguing questions of ecology and evolution (Sutherland *et al.* 2013). Evolutionary demography offers a wide array of explanations for the evolutionary processes that shape the diversity of demographic responses to environmental stochasticity (Charlesworth 1994; Healy *et al.* 2019; Hilde *et al.* 2020; Pfister 1998; Tuljapurkar *et al.* 2009). The Demographic Buffering Hypothesis (DBH, hereafter) (Morris & Doak 2004; Pélabon *et al.* 2020) is based on the fact that long-term stochastic population growth rate can be expressed in terms of the geometric mean of the annual population growth rates (Tuljapurkar 1982). As the geometric mean of λ increases, so does the long-term stochastic population growth rate (λ_s , hereafter). However, increases in variance of λ decrease λ_s (Morris & Doak 2004; Tuljapurkar 1982), and thus population persistence. This theoretical context sets the stage for the DBH. The DBH predicts that life histories are under selection pressure to minimise the negative impacts of environmental variation by constraining the temporal variance of those demographic processes (*e.g.*, survival, development, reproduction) to which population growth rate (*i.e.*, fitness) is most sensitive to (Gaillard & Yoccoz 2003; Pfister 1998). The demographic pattern operating the DBH, *i.e.*, demographic buffering, describes the selection-driven constraint on the temporal variance of the most impacting demographic processes for the population growth rate (Hilde *et al.* 2020; Morris & Doak 2004; Pfister 1998). Here, we focus on the latter - on the emerging pattern of demographic buffering in different animal life histories – rather than on the DBH itself.

A unified approach to unambiguously quantify demographic buffering is still missing. Indeed, identifying demographic buffering remains challenging (Doak *et al.* 2005; Morris & Doak 2004) for at least three reasons. First is the different interpretation of results from

correlational analyses (e.g., as in Pfister, 1998). Some authors have used the correlation coefficient as an index to order species' life histories in a continuum ranging from buffered (Spearman's correlation $\rho = <0$ between the sensitivity of λ to demographic processes and their temporal variance) to labile ($\rho = >0$), regardless of the fit of the linear regression to more or less scattered data (McDonald *et al.* 2017). In contrast, other researchers interpret the absence of statistical support for demographic buffering as an alternative strategy where variance in demographic process(es) is favoured to track environmental conditions (the so-called Demographic Lability Hypothesis, *DLH* (e.g., Koons *et al.* 2009; Reed & Slade 2012; Jäkäläniemi *et al.* 2013; Hilde *et al.* 2020). However, the increased temporal variance is a necessary but not sufficient condition to constitute demographic lability – the increased temporal variance needs to lead to (often high) change in the demographic process mean value (Le Coeur *et al.* 2022).

The second obstacle to obtain generalisation across species' populations regarding demographic buffering is the hierarchical level at which this phenomenon is typically examined. Some studies base their investigations of demographic buffering on a characteristic drawn from *the entire population model (between-populations level*, hereafter) (McDonald *et al.* 2017; Reed & Slade 2012). At the between-populations level, a life history is referred to as demographically buffered if the most important demographic process(es) has(ve) low temporal variance (Le Coeur *et al.* 2022; Hilde *et al.* 2020; Morris & Doak 2004; Pfister 1998). However, to understand how, why, and where demographic buffering occurs – or not – and how buffering patterns might be modified in response to the environment, it is essential to also consider the characteristics of the separate *components of population model (within-populations level*, hereafter). Within a population, a given demographic process can be buffered against while another can be labile to the environment (Barraquand & Yoccoz 2013; Jongejans *et al.* 2010; Koons *et al.* 2009). Thus far, studies have focused on either one

of the hierarchical levels, however, for a mechanistic understanding of how environmental stochasticity shapes life histories, both between- and within-population levels need to be addressed at the same time.

The third reason limiting a holistic understanding of demographic strategies in stochastic environments are the challenges inherent to examining their underlying mechanisms. Evidence for demographic buffering exists across some long-lived organisms with complex life cycles, (Doak *et al.* 2005; Gaillard & Yoccoz 2003; McDonald *et al.* 2017; Pfister 1998; Rotella *et al.* 2012), but also in short-lived species (Ferreira *et al.* 2013; Pfister 1998; Reed & Slade 2012). Importantly, these patterns of variation do not inform entirely on how the life histories were shaped by natural selection. The beforementioned patterns of variation are represented by *first-order* effects that perturbations in demographic processes cause on the population growth rate (*i.e.*, elasticities). A first order effect informs us regarding the population growth rate's sensitivity to *variation in demographic processes*. While a second-order effect of perturbations in demographic processes reveals the population growth rate's sensitivity to *autocorrelation* (Tuljapurkar 1990). Given so, integrating both, first and second-order effects of perturbations in demographic processes on the population growth rate, allows us to understand the behaviour of the fitness function at the vicinity of the local maxima and/or minima.

When the relationship between fitness and a demographic process is linear, the second-order derivatives of population growth rate with respect to demographic processes equal zero. In such cases, natural selection acts on the mean value of a demographic process (Shyu & Caswell 2014). However, nonzero second derivatives indicate a nonlinear relationship between fitness and a demographic process (either concave if <0 , or convex if >0) and thus provide additional and often overlooked characteristics of selection acting on

demographic processes – not only their mean values, but also their variances and covariances (Brodie *et al.* 1995; Carslake *et al.* 2008; Shyu & Caswell 2014).

The sign (*i.e.*, >0 , $=0$, <0) of the self-second derivative of λ with respect to demographic processes determines the type of (non)linear selection acting on a demographic process. For instance, a negative self-second derivative for a given demographic process describes a concave form of selection, commonly referred to as the \cap -shaped selection (Caswell 1996, 2001; Shyu & Caswell 2014). This form of selection reduces the temporal variance in said demographic process, thereby providing evidence of demographic buffering. Conversely, a demographic process yielding a positive self-second derivative identifies a convex, or U-shaped selection (Caswell 1996, 2001; Shyu & Caswell 2014). Such a selection mechanism acts upon demographic processes amplifying their temporal variance, thus potentially evidencing demographic lability (Le Coeur *et al.* 2022; Koons *et al.* 2009). The evidence of lability is only potential, because to constitute demographic lability, the increased variance needs to shift the mean value of a demographic process. The shift in the mean value of a demographic process needs to overweight the negative effect of variance on the population growth rate (Le Coeur *et al.* 2022).

The rich variation in demographic strategies across the Tree of Life is a result of evolutionary processes that have shaped variance in demographic processes through time. In this context, setting demographic buffering into the adaptive landscape context of linear and nonlinear selection enables us to identify and quantify the evolutionary processes that generate said demographic patterns. In this way, one will better understand how increased variability of environmental conditions might act on the existing –and shape novel– demographic strategies. However, we still lack a unified approach to constitute the signatures of demographic buffering.

Here, we present a framework that quantifies demographic buffering. Our framework provides a rich insight into the patterns of temporal variance in demographic processes affected by environmental stochasticity. This framework involves categorizing species or populations along a variance continuum based on the extent to which key demographic processes are buffered by natural selection, thereby limiting their temporal variability. The framework consists of four steps with a mix of well-known methods applied to stage-structured demographic information (*e.g.*, matrix population models [Caswell 2001]; integral projection models [Easterling et al. 2000]). First, we position species or populations on the aforementioned continuum to assess the cumulative effect of the variance in their key demographic processes on population growth rate at the between-populations level (see below). Second, we investigate the presence of linear selection forces operating within the life cycle of each species or population at the within-populations level (below). Third, we explore the impact of non-linear selection forces acting within the life cycle of each species or population, also at the within-populations level. The combination of these three steps provides quantitative evidence for the occurrence of demographic buffering. Step four suggests the further necessary analyses to identify demographic lability.

To demonstrate the applicability of our framework, we apply it to 40 populations of 34 mammal species sourced from the COMADRE database (Salguero-Gómez *et al.* 2016). We showcase how the framework can provide valuable insights into the patterns of demographic buffering across species. The framework offers novel, detailed insights into the selection pressures that act *within* species' life cycles, thus allowing for a thorough understanding of the evolutionary selection forces that shape the patterns of demographic buffering across species. Beyond providing a quantitative, systematic toolset to quantify buffering through three steps, we have also offer an alternative fourth step that briefly outlines how to evidence lability.

A unified framework to assess evidence of demographic buffering

The evidence for demographic buffering has been mainly assessed using Matrix Population Models (MPM; Pfister 1998; Rotella et al. 2012)). However, Integral Projection Models (IPM; Rodríguez-Caro et al. 2020; Wang et al. 2023) can be equally applied for identifying the demographic buffering signatures. Both MPMs and IPMs are stage-structured, discrete-time demographic models (Caswell 2001; Ellner *et al.* 2016). For simplicity, here we focus on MPMs, but note that the same approaches are as equally applicable to IPMs (Doak *et al.* 2021; Griffith 2017). Throughout this manuscript, we refer to demographic processes as both matrix entries a_{ij} (*i.e.*, upper-level parameters) and the vital rates that underlie the matrix elements (*i.e.*, lower-level parameters), and note that their conversion is straightforward and described elsewhere (Franco & Silvertown 2004). The framework operates on three steps.

The first step of our framework involves acquiring the relative impact of variation in demographic processes on the stochastic growth rate, λ_s , the so-called stochastic elasticities, E_{ij}^S (Haridas & Tuljapurkar 2005) (Figure 1A). The sum of all stochastic elasticities ($\Sigma E_{a_{ij}}^S$), can be separated into two components to assess how temporal variance and mean values of each demographic process impact λ_s . The first component represents the *sum of stochastic elasticity of λ_s with respect to the variance* $\Sigma E_{a_{ij}}^{S^\sigma}$, and the second represents the *sum of stochastic elasticity of λ_s with respect to the mean* $\Sigma E_{a_{ij}}^{S^\mu}$, where $\Sigma E_{a_{ij}}^S = \Sigma E_{a_{ij}}^{S^\sigma} + \Sigma E_{a_{ij}}^{S^\mu}$ (Haridas & Tuljapurkar 2005). Thus, the summation $\Sigma E_{a_{ij}}^{S^\sigma}$ quantifies the summed effect to which the stochastic population growth rate (λ_s) is influenced by changes in the variances of the demographic processes within the population matrix.

A higher sum of stochastic elasticity of λ_s with respect to the variance of demographic processes (*i.e.*, higher absolute value; $|\Sigma E_{a_{ij}}^{S\sigma}|$) indicates that small changes in the variance of demographic processes would have a substantial impact on λ_s . In other words, the variance of that demographic process is not constrained by selection, indicating absence of demographic buffering. On the other hand, a lower (absolute) stochastic elasticity of λ_s with respect to the variance of a given demographic process suggests that λ_s is less sensitive to such perturbations, or, that variance of such demographic process is being constrained by natural selection, thus pointing to demographic buffering (Haridas & Tuljapurkar 2005; Tuljapurkar *et al.* 2003) (Fig. 1A).

The first step of the framework thus features the between-populations level and places species or populations alongside a continuum. Species exhibiting unconstrained variance in demographic processes (*i.e.*, possibly not buffered, Fig. 1A, blue dots) are positioned on the left-hand side of the continuum. In contrast, species with constrained variance in demographic processes (*i.e.*, possibly buffered, Fig. 1A, yellow dots) are positioned on the right-hand side of the continuum. However, the left-hand side of the continuum does not necessarily imply evidence of demographic lability. This is so because demographic lability is defined as an increase in the *mean value* of a demographic process in response to improved environmental conditions (Le Coeur *et al.* 2022). By examining $\Sigma E_{a_{ij}}^{S\sigma}$, we can visualize an increase or decrease of the contribution that *variance* of demographic processes has on the long-term population growth rate, while the mean value of a demographic process does not change.

Step 1 of our framework examines the impacts that environmental variation has on the long-term population growth rate, λ_s (Tuljapurkar *et al.* 2003). This means that the resulting variance continuum in this step of the framework is based on how λ_s was affected by variation in the key demographic parameter across all contiguous time periods. However,

Haridas & Tuljapurkar (2005) explicitly acknowledge that covariances between demographic processes and serial correlations need to be investigated to diagnose buffering entirely. Our approach does not use covariances neither serial correlation, but rather focuses on the second derivatives of the population growth rate with respect to demographic processes and elucidates how selection is acting on variance (step 3, below).

Steps 2 and 3 of the framework are conducted at the within-populations level. Once species or populations are positioned along the variance continuum regarding the summed effect of variation on λ_s , (step 1), one needs to zoom into each life cycle separately, analysing the selection pressures acting on each one of the demographic processes composing the life cycle (*i.e.*, population model). In doing so, one can inspect the selection pressures that have generated the patterns found in step 1. Step 2 (Fig. 1B) requires obtaining the partial derivatives of the deterministic population growth rate, λ_t , relative to all matrix elements of the MPM of interest (*i.e.*, elasticities of λ_t w.r.t each demographic process in the MPM). Step 2 therefore informs on the impact that each of the demographic processes has on λ_t .

Finally, in step 3, one assesses the pattern of nonlinear selection by using the self-second derivatives of λ_t with respect to each demographic process (Fig. 1C). This step reveals the potential nonlinear selection pressures on each of the demographic processes within a life cycle. This step is key to understanding the evolutionary processes (*i.e.*, types of nonlinear selection) that the demographic processes are subjected to. Without understanding the evolutionary processes operating on the demographic processes, the pattern observed in step 1 might be artefactual. (e.g., Lawler et al. 2009)

Steps 2 and 3 of the framework feature selection pressures that have been averaged over the contiguous time periods. This means that the resulting patterns are based on how λ_t (obtained from averaging all sequential MPMs across the duration of the study) would be affected if a demographic process were perturbed. Therefore, steps 2 and 3 are based on a

different information than step 1 and can thus complete our understanding of the role of selection pressures on shaping demographic patterns across multiple species.

Another important asset of step 3 above includes the notion that the relative importance (elasticity) of demographic processes themselves changes with changing environment (Stearns 1992). In other words, the extent to which λ_t is sensitive to perturbations in a specific demographic process is *dynamic* (Kroon *et al.* 2000). Thus, the self-second derivatives generate information on how the sensitivity (or elasticity) of λ_t might change. If the sensitivity (or elasticity) of λ_t can change, then it is important to know which demographic processes are most prone to trigger such a change. In the example of a hypothetical wolf species (Fig. 1), this means that if the reproduction of the third age-class individuals (matrix element $a_{1,3}$) decreased, the sensitivity of λ_t to $a_{1,3}$ would increase (square with the largest black dot, Fig. 1C). Consequently, with increased environmental variability, the key demographic process might change from remaining in the fourth age class (matrix element $a_{4,4}$, Fig. 1B) to reproduction of the third age-class (matrix element $a_{1,3}$, Fig. 1C).

Combining the three steps of our framework allows for a quantitative identification of buffering. Steps 2 and 3 offer key insights as to *why* a given species or population is placed on either the buffered or the non-buffered end of the variance continuum. A clear and unequivocal evidence for support towards buffering consists of: (1) a species or population being positioned near the 0 end of the continuum (the right-hand side) in step 1; (2) this species' or populations' life cycle having one or more demographic processes with highest elasticity values in step 2; and (3) the same demographic process displaying the highest elasticity in step 2 with negative self-second derivative values in step 3. In this sense, Figure 1B shows that, for the chosen population of a hypothetical wolf species, the most important demographic process is remaining in the fourth stage (MPM element $a_{4,4}$), as this demographic process results in highest elasticity value (Fig. 1B yellow square). However,

Fig. 1C reveals that $a_{4,4}$ is under little selection pressure for variance reduction. Thus, there is no clear evidence of buffering from the third step of the framework (*i.e.*, no concave selection forces). This way, the lack of concave selection forces on the key demographic process within wolf's life cycle explains why this species is placed on the left-hand side of the variance continuum (Fig. 1A).

Species placed on the non-buffered end of the continuum is a necessary but not sufficient condition for evidence demographic lability. It is key highlighting here that demographic buffering and lability do not represent two extremes of the same continuum. The variance continuum allocates the species or populations from strongly buffered to non-buffered, but to test the for lability, a further step is needed.

Although not our primary goal here, we briefly introduce said step 4. To establish compelling evidence of lability, it is essential to fulfil several further criteria. First, sufficient data across various environments (over time or space) are required to construct reaction norms that depict how a demographic process responds to environmental changes (Morris et al., 2008; Koons et al., 2009), which can be challenging in terms of sufficient and high-quality demographic and environmental data. Second, non-linear relationships between demographic processes and the environment must be established based on the demographic process-environment reaction norms. Lastly, demographic processes where an increase in the mean value has a stronger positive impact on population growth rate than the detrimental effect of increased variance needs to be identified. The latter condition is only met when the demographic process-environment reaction norm takes a convex shape (resembling a "U" shape), as described by Koons et al. (2009) and Morris et al. (2008). However, a study by Barraquand and Yoccoz (2013) reported diverging results in this regard. Importantly, we note that more likely than previously thought (*e.g.*, Pfister 1998), species do not exist as purely buffering or labile, but that within populations, some vital rates may be buffered, other labile,

and others insensitive to the environment (*e.g.*, (Doak *et al.* 2005). Deciphering generality in this likely complex pattern should attract much research attention going forward, in our opinion.

Demographic buffering in mammals: a case study using the unified framework

We demonstrate the performance of our framework using 44 MPMs from 34 mammal species. Mammals are of special interest here for two reasons: (1) mammalian life histories have been well studied (Bielby *et al.* 2007; Gillespie 1977; Jones 2011; Stearns 1983); and (2) some of their populations have already been assessed in terms of buffering, particularly for primates (Campos *et al.* 2017; Morris *et al.* 2008, 2011; Reed & Slade 2012; Rotella *et al.* 2012). Together, the well-studied life histories and previous information about the occurrence of buffering in mammals provide the necessary information to make accurate predictions and validate the performance of the proposed framework.

We used Matrix Population Models from 40 out of 139 studies with mammals available in the COMADRE database v.3.0.0 (Salguero-Gómez *et al.* 2016). These 40 populations encompass 34 species from eight taxonomic orders. We included these MPMs in our analyses because they provide values of demographic processes (a_{ij}) for three or more contiguous time periods, thus allowing us to obtain the stochastic elasticity of each a_{ij} . Although we are aware that not all possible temporal variation in demographic processes may have been expressed within this period, we assumed three or more transitions are enough to provide sufficient variation for population comparison. At least three contiguous time periods - a common selection criteria in comparative studies of stochastic demography (Compagnoni *et al.* 2023) - also allowed to test and showcase our framework. Fortunately, several long-lived species, characterized by low variation in their demographic processes, were studied for a long time (*e.g.*, some primates in our dataset have been studied for over 20 years – Morris *et al.* 2011). We removed the populations where either only survival or only reproduction

rates were reported, because of the impossibility to calculate the stochastic growth rate. A detailed description of the analysed data and their original sources are available in supplementary material (Supplementary Material, Table S1).

Homo sapiens was included in our analyses because it is the only mammalian species in which second-order derivatives have been applied (Caswell 1996). Therefore, *Homo sapiens* provides an ideal basis for comparisons among species. The data for *Homo sapiens* were gathered from 26 modern populations located in various cities, allowing us to construct a spatiotemporal variance. It is important to note that in this case, we are not working with true temporal variance but rather a variance that encompasses both spatial and temporal aspects.

For steps 2 and 3 of our framework, we utilized a subset of 16 populations (including *Homo sapiens*) whose population projection matrices (MPMs) were organized by age. We specifically selected these populations because their life cycles can be summarized by two main demographic processes: survival and contribution to recruitment of new individuals. The contribution to recruitment can be interpreted as either the mean reproductive output for each age class or an approximation thereof, depending on how the matrices are structured (Ebert 1999). One advantage of using such matrices is that they encompass only two types of demographic processes, namely survival and recruitment, eliminating the need to account for multiple transitions between different life stages.

To perform the step 1 of our framework and obtain the $\Sigma E_{a_{ij}}^{S^\sigma}$ (and $\Sigma E_{a_{ij}}^{S^\mu}$), we followed Tuljapurkar *et al.* (2003) and Haridas & Tuljapurkar (2005). To perform step 2 of our framework, we calculated the deterministic elasticities of each demographic process extracted using the *popbio* package. All analyses were performed using R version 3.5.1 (R Core team, 2018). Finally, to perform the step 3 of our framework the self-second derivatives were

adapted from *demogR* (Jones 2007) following (Caswell 1996) and applied for the mean MPM.

Results

We ranked 40 populations from the 34 identified mammal species according to the cumulative impact of variation in demographic processes on λ_s using the step 1 of our framework (Fig. 2). Additional information (including standard deviations of the elasticity estimates and number of matrices available) is provided in the supplementary material (Table S1). Most of the analysed orders were placed on the low-variance end of the variance continuum (Fig. 2). The smallest contributions of variation in demographic processes (*i.e.*, maximum value of $\Sigma E_{a_{ij}}^{S^\sigma}$, note that $\Sigma E_{a_{ij}}^{S^\sigma}$ ranges from 0 to -1), suggesting more buffered populations, were assigned to Primates: northern muriqui (*Brachyteles hyphoxantus*, $\Sigma E_{a_{ij}}^{S^\sigma} = -0.09 \times 10^{-4} \pm 0.12 \times 10^{-4}$) (mean \pm standard deviation) (Fig. 2 silhouette a), mountain gorilla (*Gorilla beringhei*, $\Sigma E_{a_{ij}}^{S^\sigma} = -0.24 \times 10^{-4} \pm 0.08 \times 10^{-4}$) (Fig. 2 silhouette b), followed by the blue monkey (*Cercopithecus mitis*, $\Sigma E_{a_{ij}}^{S^\sigma} = -0.63 \times 10^{-4} \pm 0.06 \times 10^{-4}$) (Fig. 2 silhouette c). The first non-primate species placed near the low-variance end of the continuum was the Columbian ground squirrel (*Urocitellus columbianus*, Rodentia, $\Sigma E_{a_{ij}}^{S^\sigma} = -0.003 \pm 0.002$) (Fig. 2 silhouette d). The species with the highest contribution of variation in demographic processes placed at the high-variance end of the continuum was the stoat (*Mustela erminea*, Carnivora, $\Sigma E_{a_{ij}}^{S^\sigma} = -0.35 \pm 0.02$) (Fig. 2 silhouette e). All the 14 primate populations displayed potential evidence of buffering, occupying the right-hand side of the variance continuum, with the exception of the Patas monkey (*Erythrocebus patas*, Primates, $\Sigma E_{a_{ij}}^{S^\sigma} = -0.05 \pm 0.03$) (Fig. 2 silhouette f). The snowshoe hare (*Lepus americanus*, Lagomorpha, $\Sigma E_{a_{ij}}^{S^\sigma}$

385 = -0.29 ± 0.16) (Fig. 2 silhouette g) and the Bush rat (*Rattus fuscipes*, Rodentia, $\Sigma E_{a_{ij}}^{S\sigma} = -0.25$
386 ± 0.03) (Fig. 2 silhouette h) appear on the high-variance end of the continuum.

387 As predicted for the steps 2 and 3, we could not observe a clear pattern in support of
388 buffering. This finding means that the demographic processes with the highest elasticity
389 values failed to display strongly negative self-second derivatives (Fig. 3). Particularly for
390 majority of primates - with the lack or minor temporal variation in demographic processes -
391 demographic processes with high elasticities had positive values for the self-second
392 derivatives (indicated by yellow squares with white dots in Fig. 3). Examples of primate
393 species exhibiting high elasticities and positive values for the self-second derivatives and
394 include northern muriqui (*Brachyteles hypoxanthus*), mountain gorilla (*Gorilla beringei*),
395 white-faced capuchin monkey (*Cebus capucinus*), rhesus monkey (*Macaca mulatta*), blue
396 monkey (*Cercopithecus mitis*), Verreaux's sifaka (*Propithecus verreauxi*) and olive baboon
397 (*Papio cynocephalus*) (Fig. 3). This implies that the key demographic processes influencing
398 λ_t are not subject to selective pressure for reducing their temporal variability. However, even
399 though the primates were positioned closer to the low-variance end of the continuum in step
400 1, the evidence from steps 2 and 3 does not support the occurrence of buffering in the most
401 influential demographic processes.

402 The killer whale showed similar controversy between step 1 and steps 2-3 results as
403 most primates. In step 1, the killer whale was positioned at the buffered end of the variance
404 continuum (*Orcinus orca*, Cetacea, $\Sigma E_{a_{ij}}^{S\sigma} = -0.70 \times 10^{-4} \pm 1.04 \times 10^{-5}$) (Fig. 2 silhouette not
405 shown). However, steps 2 and 3 show that the three demographic processes in killer whale
406 life cycle with highest elasticity values (matrix elements $a_{2,2}$, $a_{3,3}$ and $a_{4,4}$) are not under
407 selection pressures for reducing their temporal variance, but the opposite (depicted by yellow
408 and green squares with white dots, Fig. 3).

The only primate species exhibiting evidence of buffering in steps 2 and 3 was human. In human, demographic parameters representing survival from first to second age class (matrix element $a_{2,1}$) displayed high elasticities and negative self-second derivatives (depicted as yellow squares with black dots in Fig. 3). Evidence supporting buffering was also found in the Columbian ground squirrel (*Urocitellus columbianus*), where, similar to humans, survival from the first to the second age class (matrix element $a_{2,1}$) showed indications of selection acting to reduce $a_{2,1}$ variance. Accordingly, the Columbian ground squirrel was positioned close to the buffered end of the variance continuum in step 1. Hence, the Columbian ground squirrel was the sole species with consistent evidence of buffering across all three steps of the framework.

The Soay sheep (*Ovis aries*) was the species furthest from the buffered end of the variance continuum that enabled to perform steps 2 and 3. For the Soay sheep, remaining in the third age class (matrix element $a_{3,3}$) has the major influence on λ_t and is under selection pressure to have its variance increased. The latter characteristics reveal potential conditions for lability even though the species is placed closer to the buffered end of the variance continuum.

Steps 2 and 3 illustrate the importance of examining buffering evidence on the within-populations level. These two steps of the framework identify the simultaneous acting of concave and convex selection on different demographic processes within a single life cycle. In polar bear (*Ursus maritimus*), the key demographic process (matrix element $a_{4,4}$) is under convex selection, as depicted by a yellow square with a white dot in Fig. 3. However, the demographic process with the second highest elasticity value (matrix element $a_{5,4}$) is under strong concave selection (depicted by a light green square with a black dot in Fig. 3).

By adding step 3 to the framework, another important information was accessed. The high absolute values of self-second derivatives (large dots, either black or white, Fig. 3) indicate where the sensitivity of λ_t to demographic parameters is itself prone to environmental changes. For instance, if the value of $a_{5,4}$ for polar bear increased, the sensitivity of λ_t to $a_{5,4}$ would decrease because the self-second derivative of $a_{5,4}$ is highly negative (depicted by the largest black dot in polar bear MPM). Vice versa holds for the $a_{4,4}$ demographic process, where an increase in the value of $a_{4,4}$ would increase λ_t 's sensitivity to $a_{4,4}$, because the self-second derivative of $a_{5,4}$ is highly positive (depicted by the largest white dot in polar bear MPM). Thus, sensitivities (or equally elasticities) of demographic processes with high absolute values for self-second derivatives can easily change..

Discussion

In the Anthropocene, identifying and quantifying mechanisms of species responses to stochastic environments holds crucial importance. This importance is particularly tangible in the context of the unprecedented environmental changes and uncertainties that impact the dynamics and persistence of natural populations (Boyce *et al.* 2006). Correlational demographic analysis, whereby the importance of demographic processes and their temporal variability is examined (Pfister 1998), has attempted to identify how species may buffer against the negative effects of environmental stochasticity. However, these widely used approaches have important limitations (see Introduction and Hilde *et al.* 2020). One significant limitation is the issue of measurement scale concerning demographic processes (Hilde *et al.* 2020; Morris & Doak 2004). Demographic processes, such as birth rates, death rates, immigration, and emigration, operate at various temporal and spatial scales. The choice of scale at which these processes are measured can impact the outcomes of correlational demographic analysis (Bjørkvoll *et al.* 2016). Our novel framework overcomes said

limitations by providing a rigorous approach to quantify demographic buffering (Hilde et al. 2020; Pfister 1998).

Evidencing demographic buffering is not straightforward. Indeed, through the analysis of stochastic population growth rate (λ_s) in our application of the framework to 44 populations of 34 species, we identify the highest density of natural populations near the buffered end of the variance continuum (step 1). However, we show that the same species then fail to exhibit signs of concave (\cap -shaped) selection on the key demographic parameters when further analyses are performed averaging the variation across the duration of each study (steps 2 and 3). This finding confirms that placing the species near the buffered end of the variance continuum is *necessary* but not *sufficient* to diagnose demographic buffering. Indeed, buffering occurs when concave selection forces act on the key demographic parameter (Caswell 1996, 2001; Shyu & Caswell 2014).

Combining the three steps into a unified framework is of outmost importance. In steps 2 and 3 of the framework, we find relatively limited overall evidence of buffering in the examination of our 16 (out of 34 in step 1) studied animal species. Step 3 of our framework reveals that the role of natural selection shaping temporal variation in demographic processes is more complex than expected. Indeed, demographic processes within our study populations are often under a mix of convex and concave selection. This mix of selection patterns was already suggested by Doak *et al.* (2005). Here, only two out of 16 mammal species revealed concave selection acting on the key demographic processes (Columbian ground squirrel [*Urocitellus columbianus*], and humans, [*Homo sapiens sapiens*]). These two species were also placed near the buffered end of the variance continuum, therefore meeting all the necessary conditions to diagnose buffering. However, finding 12.5% (two out of 16) species that meet the criteria for demographic buffering is not in concordance with previous studies. Evidence of buffering has been reported across 22 ungulate species (Gaillard & Yoccoz

2003). In the one ungulate we examined, the moose (*Alces alces*), we find only partial support for buffering in adult survival, since this species is placed near the buffered end of the variance continuum in step 1 but does not show concave selection pressures on adult survival in step 2/3, as would be necessary to confirm the occurrence of buffering. It is worth noting that a varying number of matrices per species were employed, ranging from 1 to 21, with an average of 8.1 matrices per species (as shown in Table S1). Naturally, having a greater number of matrices is preferred in such analyses. Furthermore, while the size of matrices (matrix dimensions) does not directly bias the results of our framework in any way – since steps 2 and 3 are shown for all the demographic processes independent of matrix dimension – potential implications of varying matrix dimensions should be further investigated in the future.

Our overall findings reveal varying levels of support for the notion that adult survival in long-lived species tends to be buffered. Indeed, Gaillard *et al.* (1998) found that adult female survival varied considerably less than juvenile survival in large herbivores. This finding was also supported by further studies in ungulates and small rodents (Gaillard & Yoccoz 2003), turtles (Heppell 1998), vertebrates and plants (Pfister 1998), and more recently across nine (out of 73) species of plants (McDonald *et al.* 2017).

When placing our study species along a variance continuum (step 1), primates tend to be located on the buffered end. However, most primates displayed convex –instead of the expected concave– selection on adult survival. Similar results, where the key demographic process failed to display constrained temporal variability, have been reported for long-lived seabirds (Doherty *et al.* 2004). One explanation for the unexpected convex selection on adult survival involves trade-offs, as suggested by (Doak *et al.* 2005). When two demographic parameters are negatively correlated, the variance of population growth rate (λ) can be increased or decreased (Compagnoni *et al.* 2016; Evans & Holsinger 2012). The well-

established trade-off between survival and fecundity (Roff & Fairbairn 2007; Stearns 1992) might explain the observed deviation of our results. Because variation in primate recruitment is already constrained by physiological limitations (Campos *et al.* 2017), when adult survival and recruitment are engaged in a trade-off, this trade-off might lead to our unexpected result. Correlations among demographic processes (positive and negative) inherently influence the biological limits of variance (Haridas & Tuljapurkar, 2005). This is because the magnitude of variation in a particular demographic process is constrained by (the variation of) other demographic processes that exert an influence on it. Not surprisingly, correlations among demographic processes have been shown to be strongly subjected to ecological factors (Fay *et al.* 2022). Here, future studies may benefit from deeper insights via cross-second derivatives (Caswell 1996, 2001) to investigate correlations among demographic processes.

Examining the drivers of demographic buffering has become an important piece of the ecological and evolutionary puzzle of demography. As such, quantifying buffering can help us better predict population responses to environmental variability, climate change, and direct anthropogenic disturbances (Boyce *et al.* 2006; McDonald *et al.* 2017; Pfister 1998; Vázquez *et al.* 2017). By setting demographic buffering into a broader and integrated framework, we hope to enhance comprehension and prediction of the implications of heightened environmental stochasticity on the evolution of life history traits. This understanding is crucial in mitigating the risk of extinction for the most vulnerable species.

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Data availability

The demographic data used in this paper are open-access and available in the COMADRE Animal Matrix Database (<https://compadre-db.org/Data/Comadre>). A list of the studies and species used here is available in Supplementary Material (Table S1). The data and code supporting the results can be accessed here: https://github.com/SamuelGascoigne/Demographic_buffering_unified_framework.

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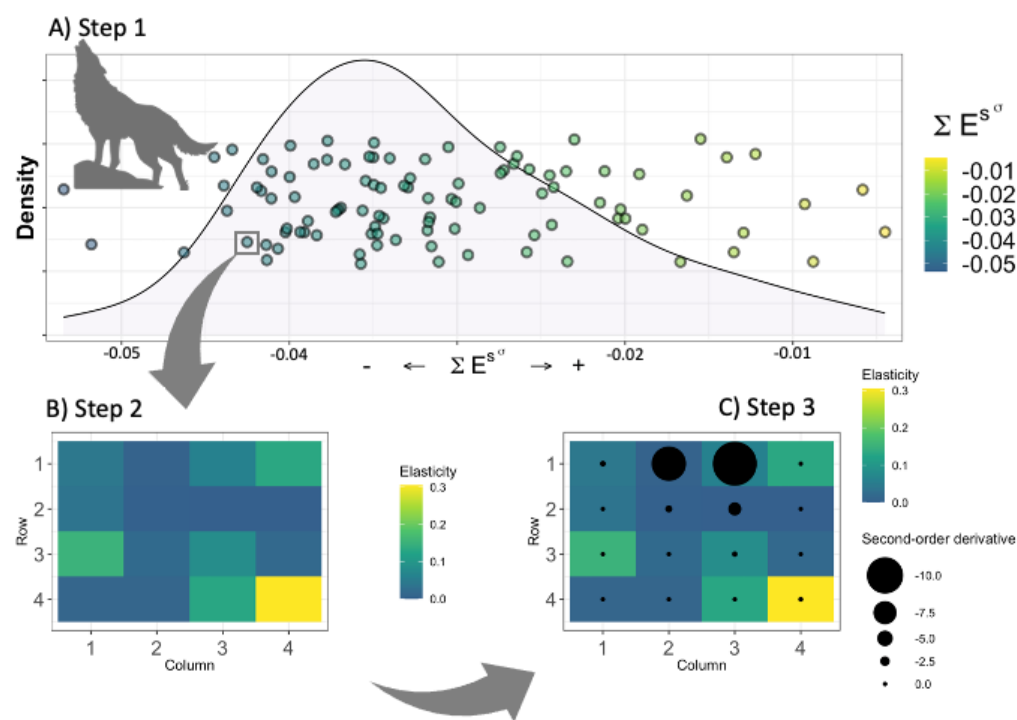
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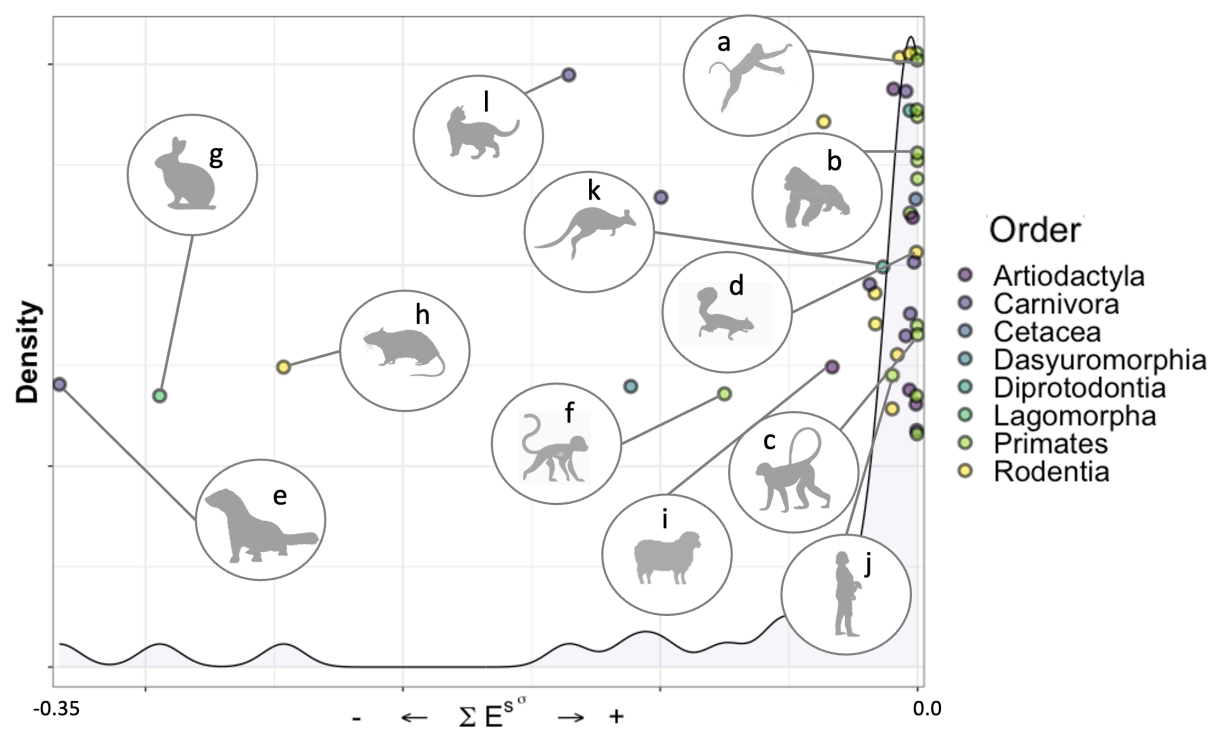
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684



687 **Figure 2**



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690 **Figure 3**

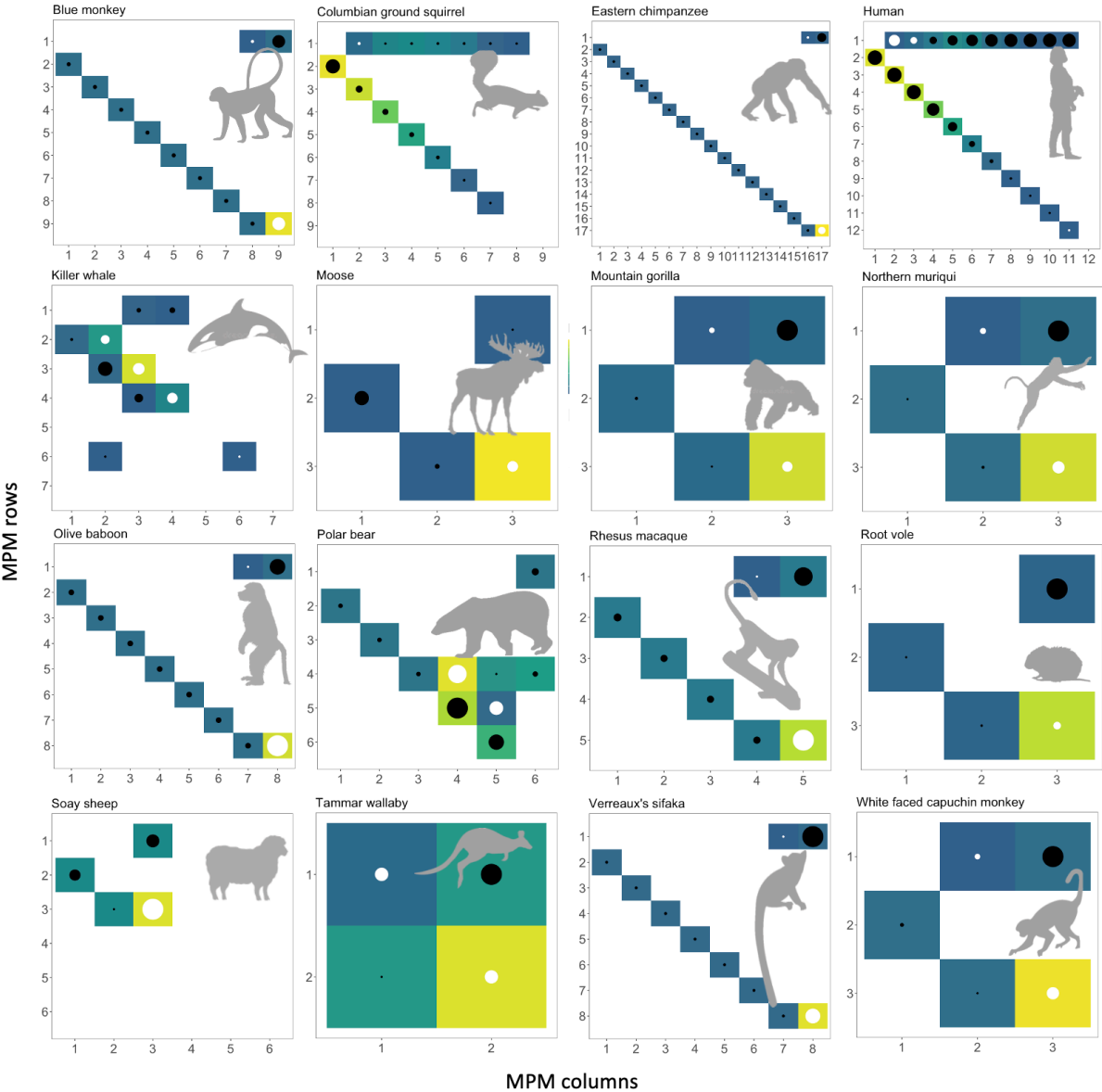


Figure legends

Figure 1. A three-step framework proposed to: Step 1 - allocate species and/or populations on a variance continuum (plot A, dots representing 50 hypothetical species). The variance continuum operates at the between-populations level (see text) and is represented by partitioning the sum of all the stochastic elasticities ($\Sigma E_{a_{ij}}^S$) into two compounds: i) sums of stochastic elasticities with respect to the variance ($\Sigma E_{a_{ij}}^{S^\sigma}$), and ii) sums of stochastic elasticities with respect to the mean ($\Sigma E_{a_{ij}}^{S^\mu}$). The first step of our framework shows the variance compound of the sums of stochastic elasticities forming a continuum where the right-hand side of the plot represents species (or populations) where a perturbation of variance in demographic processes results in weak or no impact on λ_s (yellow dots). The yellow-dotted species (or populations) can be classified as having potentially *buffered life-cycles* – based on all the demographic processes. The left-hand side of the graph represents species (or populations) where a perturbation of the variance in demographic processes results in strong impact on λ_s (blue dots). Thus, the blue-dotted species (or populations) can be classified as having potentially *unbuffered life cycles* – based on all the demographic processes. The vertical axis delineates the values of the density distribution function, indicating the number of species/populations at each value of $\Sigma E_{a_{ij}}^{S^\sigma}$. The placement of data points (species/populations) along the horizontal axis corresponds to their calculated values of $\Sigma E_{a_{ij}}^{S^\sigma}$ and is arranged linearly, while the breadth along the y-axis is solely for improved visual comprehension. Step 2 - Access the linear selection pressures for individual species or populations at within-species level (see text) (plot B). Step 2 displays the elasticities of the deterministic population growth rate (λ_t) for a hypothetical population of wolf and reveals the linear selection gradients, and which demographic processes are the most influential for λ_t . Step 3 - Access the nonlinear selection pressures at the within-species level (see text) (plot

C). In the third step self-second derivatives for the corresponding demographic processes from step 2 are displayed.

Figure 2. Results for step 1 of our framework showing the sum of stochastic elasticities with respect to the variance $\Sigma E_{a_{ij}}^{S\sigma}$. The closer the $\Sigma E_{a_{ij}}^{S\sigma}$ is to zero, the weaker the impact of variation in demographic processes on λ_s . The 40 populations from 34 species of mammals from the COMADRE database are ranked into the variance continuum from potentially buffered (right-hand side) to less buffered (left-hand side), since any variation in demographic processes would strongly impact λ_s . Colors represent different taxonomic orders with Primates occupying the right-hand side. Silhouettes: a) *Brachyteles hypoxantus*, b) *Gorilla beringhei*, c) *Cercopithecus mitis*, d) *Urocyon columbianus*, e) *Mustela erminea*, f) *Erythrocebus patas*, g) *Lepus americanus*, h) *Rattus fuscipes*, i) *Ovis aries*, j) *Homo sapiens*, k) *Macropus eugenii*, and l) *Felis catus*. The vertical axis delineates the values of the density distribution function, indicating the number of species/populations at each value of $\Sigma E_{a_{ij}}^{S\sigma}$. The placement of data points (species/populations) along the horizontal axis corresponds to their calculated values of $\Sigma E_{a_{ij}}^{S\sigma}$ and is arranged linearly, while the breadth along the y-axis is solely for improved visual comprehension.

Figure 3: Results from steps 2 and 3 of the proposed framework (see Fig. 2B, C). The 16 plots represent populations where the MPMs built by ages were available in the COMADRE database (see text). The color scale represents elasticity values for each of the demographic processes in the MPM, where yellow represents high and blue low elasticity values. No color means elasticity=0. Because the aim of step 2 is to identify the most impacting demographic process within each species' life cycle (the within-populations level, see text) - not to

743 compare the elasticity values among species - each plot has its own scale (see end of legend).
 744 The black dots represent negative self-second derivatives of λ_t - thus concave selection - and
 745 the white dots represent positive self-second derivatives of λ_t - thus convex selection. The dot
 746 sizes are scaled by the absolute value of self-second derivatives, where the smaller the dot,
 747 the closer a self-second derivative is to 0, indicating weak or no nonlinearity. Large dots
 748 indicate strong nonlinear selection forces. Scales ($E_{\min-\max}$ =elasticity minimum and maximum
 749 value, $SSD_{\min-\max}$ =self-second derivative minimum and maximum value): Blue monkey $E_{\min-$
 750 $\max=0.00-0.52$, $SSD_{\min-\max}=-1.25-1.27$; Columbian ground squirrel: $E_{\min-\max}=0.00-0.23$,
 751 $SSD_{\min-\max}=-1.48-0.01$; Eastern chimpanzee: $E_{\min-\max}=0.00-0.60$, $SSD_{\min-\max}=-4.39-2.59$;
 752 Human: $E_{\min-\max}=0.00-0.18$, $SSD_{\min-\max}=-0.15-0.08$; Killer whale: $E_{\min-\max}=0.00-0.55$,
 753 $SSD_{\min-\max}=-5.72-3.43$; Moose: $E_{\min-\max}=0.00-0.55$, $SSD_{\min-\max}=-0.66-0.36$; Mountain gorilla:
 754 $E_{\min-\max}=0.00-0.81$, $SSD_{\min-\max}=-1.46-0.28$; Northern muriqui: $E_{\min-\max}=0.00-0.72$, $SSD_{\min-$
 755 $\max=-1.17-0.35$; Olive baboon: $E_{\min-\max}=0.00-0.54$, $SSD_{\min-\max}=-0.57-1.13$; Polar bear: $E_{\min-$
 756 $\max=0.00-0.26$, $SSD_{\min-\max}=-0.73-0.54$; Rhesus macaque: $E_{\min-\max}=0.00-0.51$, $SSD_{\min-\max}=-$
 757 $0.54-0.71$; Root vole: $E_{\min-\max}=0.00-0.86$, $SSD_{\min-\max}=-2.54-0.22$; Soay sheep: $E_{\min-\max}=0.00-$
 758 0.56 , $SSD_{\min-\max}=-0.22-0.40$; Tammar wallaby: $E_{\min-\max}=0.00-0.55$, $SSD_{\min-\max}=-0.64-0.34$;
 759 White faced capuchin monkey: $E_{\min-\max}=0.00-0.66$, $SSD_{\min-\max}=-2.66-1.21$.
 760

761 **Supplementary material – Data available in COMADRE Version 2.0.1 and results from Step 1 of the framework**

762 **Table S1.** The metadata used in step 1 of our framework and the respective results presented in the main text. The first four columns represent
763 the information from where Matrix Populations Models (MPMs) were extract precisely as presented in COMADRE 2.0.1. Column titles differ
764 from the database as “SpeciesAuthorComadre” is equivalent to “SpeciesAuthor” and “SpeciesName” is equivalent to “SpeciesAccepted” in
765 COMADRE 2.0.1. The remaining columns present the results of step 1, where we present the raw values of $\Sigma E_{a_{ij}}^{S^{\mu}}$ and $\Sigma E_{a_{ij}}^{S^{\sigma}}$, their respective
766 standard deviation, the stochastic population growth rate λ_s , and the number of available matrices (# matrices). For ByAge, “TRUE” was
767 assigned for MPMs built by age or “FALSE” if otherwise.

SpeciesAuthorComadre	SpeciesName	CommonName	Order	$\Sigma E_{a_{ij}}^{S^{\mu}}$	$\Sigma E_{a_{ij}}^{S^{\mu}}$ (sd)	$\Sigma E_{a_{ij}}^{S^{\sigma}}$	$\Sigma E_{a_{ij}}^{S^{\sigma}}$ (sd)	# matrices	λ
Homo_sapiens_subsp._sapiens	<i>Homo sapiens sapiens</i>	Human	Primates	1.003	0.003	1.003	0.004	13	1.064
Alces_alces	<i>Alces alces</i>	Moose	Artiodactyla	1.001	0.001	1.001	0.001	13	1.205
Antechinus_agilis	<i>Antechinus agilis</i>	Agile antechinus	Dasyuromorphia	1.111	0.111	1.111	0.011	2	0.931
Brachyteles_hypoxanthus	<i>Brachyteles hypoxanthus</i>	Northern muriqui	Primates	1.000	0.000	1.000	0.000	12	1.051
Callospermophilus_lateralis	<i>Callospermophilus lateralis</i>	Golden-mantled ground squirrel	Rodentia	1.054	0.054	1.054	0.055	9	2.052
Cebus_capucinus	<i>Cebus capucinus</i>	White faced capuchin monkey	Primates	1.000	0.000	1.000	0.000	11	1.021
Cercopithecus_mitis	<i>Cercopithecus mitis</i>	Blue monkey	Primates	1.000	0.000	1.000	0.000	14	1.036
Eumetopias_jubatus	<i>Eumetopias jubatus</i>	Northern sea lion; Steller sea lion	Carnivora	1.005	0.005	1.005	0.002	2	0.904
Felis_catus	<i>Felis catus</i>	Feral cat	Carnivora	1.136	0.136	1.136	0.012	1	1.948
Gorilla_beringei	<i>Gorilla beringei</i>	Mountain gorilla	Primates	1.000	0.000	1.000	0.000	21	1.027

Hippocamelus_bisulcus	<i>Hippocamelus bisulcus</i>	Huemul deer	Artiodactyla	1.002	0.002	1.002	0.001	1	0.996
Lepus_americanus	<i>Lepus americanus</i>	Snowshoe hare	Lagomorpha	1.294	0.294	1.294	0.165	2	0.812
Lycaon_pictus	<i>Lycaon pictus</i>	African wild dog	Carnivora	1.100	0.100	1.100	0.008	1	1.500
Macaca_mulatta_3	<i>Macaca mulatta</i>	Rhesus macaque	Primates	1.000	0.000	1.000	0.001	12	1.127
Macropus_eugenii	<i>Macropus eugenii</i>	Tammar wallaby	Diprotodontia	1.013	0.013	1.013	0.012	7	0.981
Marmota_flaviventris_2	<i>Marmota flaviventris</i>	Yellow-bellied marmot	Rodentia	1.007	0.007	1.007	0.006	4	0.890
Marmota_flaviventris_3	<i>Marmota flaviventris</i>	Yellow-bellied marmot	Rodentia	1.008	0.008	1.008	0.005	4	0.921
Microtus_oekonomus	<i>Microtus oeconomus</i>	Root vole	Rodentia	1.000	0.000	1.000	0.001	14	1.028
Mustela_erminea	<i>Mustela erminea</i>	Stoat	Carnivora	1.334	0.334	1.334	0.117	2	1.258
Orcinus_orca_2	<i>Orcinus orca</i>	Killer whale	Cetacea	1.001	0.001	1.001	0.001	24	0.999
Ovis_aries_2	<i>Ovis aries</i>	Soay sheep	Artiodactyla	1.033	0.033	1.033	0.020	3	1.099
Pan_troglodytes_subsp._schweinfurthii	<i>Pan troglodytes</i>	Eastern chimpanzee	Primates	1.000	0.000	1.000	0.001	22	0.982
Papio_cynocephalus	<i>Papio cynocephalus</i>	Olive baboon	Primates	1.000	0.000	1.000	0.000	19	1.054
Peromyscus_maniculatus_2	<i>Peromyscus maniculatus</i>	Deer mouse	Rodentia	1.010	0.010	1.010	0.005	2	1.107
Phocarctos_hookeri	<i>Phocarctos hookeri</i>	New Zealand sea lion	Carnivora	1.005	0.005	1.005	0.003	8	1.023
Propithecus_verreauxi	<i>Propithecus verreauxi</i>	Verreaux's sifaka	Primates	1.000	0.000	1.000	0.000	12	0.986
Puma_concolor_8	<i>Puma concolor</i>	Cougar	Carnivora	NA	NA	NA	NA	10	1.115
Rattus_fuscipes	<i>Rattus fuscipes</i>	Bush rat	Rodentia	1.246	0.246	1.246	0.029	2	1.305
Spermophilus_armatus	<i>Urocitellus armatus</i>	Uinta ground squirrel	Rodentia	1.016	0.016	1.016	0.011	4	1.125
Spermophilus_armatus_2	<i>Urocitellus armatus</i>	Uinta ground squirrel	Rodentia	1.017	0.017	1.017	0.010	3	1.095

Spermophilus_columbianus	<i>Urocitellus columbianus</i>	Columbian ground squirrel	Rodentia	1.036	0.036	1.036	0.025	3	1.009
Spermophilus_columbianus_3	<i>Urocitellus columbianus</i>	Columbian ground squirrel	Rodentia	1.003	0.003	1.003	0.006	3	1.200
Ursus_americanus_subsp._floridanus	<i>Ursus americanus</i>	Florida black bear	Carnivora	1.003	0.003	1.003	0.003	2	1.020
Ursus_arctos_subsp._horribilis_5	<i>Ursus arctos</i>	Grizzly bear	Carnivora	1.001	0.001	1.001	0.001	4	1.026
Ursus_maritimus_2	<i>Ursus maritimus</i>	Polar bear	Carnivora	1.019	0.019	1.019	0.007	2	0.941
Brachyteles_hypoxanthus_2	<i>Brachyteles hypoxanthus</i>	Northern muriqui	Primates	1.000	0.000	1.000	0.000	12	1.111
Cebus_capucinus_2	<i>Cebus capucinus</i>	WhiteNAfaced capuchin monkey	Primates	1.000	0.000	1.000	0.000	11	1.059
Chlorocebus_aethiops_2	<i>Chlorocebus aethiops</i>	Vervet	Primates	1.075	0.075	1.075	0.087	5	1.187
Erythrocebus_patas	<i>Erythrocebus patas</i>	Patas monkey	Primates	1.051	0.051	1.051	0.038	5	1.128
Gorilla_beringei_subsp._beringei 768	<i>Gorilla beringei</i>	Mountain gorilla	Primates	1.000	0.000	1.000	0.000	21	1.053
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